

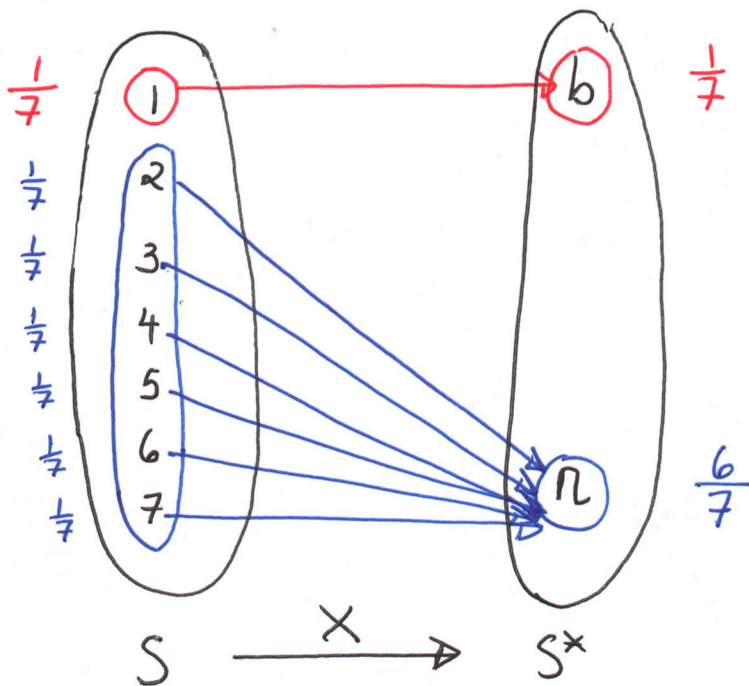
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## Random Variables Lecture 1

- Generate an induced sample space.
  - Transition from given sample space to one of more significant interest.

For example, when we play Russian roulette with 7-chamber pistol, given sample space  $S = \{1, 2, 3, 4, 5, 6, 7\}$  where outcome is chamber in line of fire.

Induced Sample space is  $S^* = \{b, n\}$  (bullet, no bullet).



- $X$  maps 1 to  $b$  and  $2, \dots, 7$  to  $n$ .
- Think of outcomes in  $S$  as pebbles, and of the probabilities of these outcomes as the weights of these stones.  $X$  crushes these stones together, adding corresponding

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weights.

Def. Let  $X: S \xrightarrow{\text{onto}} S^*$  be a surjective function from a sample space  $S$  onto a set  $S^*$ . Then  $X$  is a random variable if  $X$  induces a probability on  $S^*$

$$\text{by } P(y) = P(X^{-1}(y)) = \sum_{x \in S: X(x)=y} P(x)$$

In the previous example, this just means

$$P(b) = P(X^{-1}(b)) = P(1) = \frac{1}{7}$$

$$P(r) = P(X^{-1}(r)) = \sum_{k=2}^7 P(k) = \frac{6}{7}$$

Ex. 3 balls randomly selected without replacement from urn containing 20 balls numbered 1-20. If we bet that at least one of the chosen balls has a number as large as or larger than 17, what's the probability that we win the bet?

Solution: Let  $X = 3, 4, \dots, 20$  be the largest denomination of the 3 chosen balls. We want

$$\begin{aligned} P(X \geq 17) &= P(X=17) + P(X=18) + P(X=19) + P(X=20) \\ &= \frac{\binom{16}{2} + \binom{17}{2} + \binom{18}{2} + \binom{19}{2}}{\binom{20}{3}} \approx 0.508 \end{aligned}$$

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Ex. A fair die is tossed until it comes up 6. Let  $X$  denote the number of required tosses.

Then  $X = 1, 2, 3, \dots, n, \dots, \infty$ .

$$P(X=1) = \frac{1}{6}$$

$$P(X=2) = \frac{5}{6} \cdot \frac{1}{6}$$

$$\vdots$$
$$P(X=n) = \left(\frac{5}{6}\right)^{n-1} \frac{1}{6}$$

The probability that finitely many tosses is required

$$P(X \in \mathbb{N}) = P\left(\bigcup_{n=1}^{\infty} \{X=n\}\right) = \sum_{n=1}^{\infty} P(X=n) =$$
$$= \sum_{n=1}^{\infty} \left(\frac{5}{6}\right)^{n-1} \frac{1}{6} = \frac{1}{1-\frac{5}{6}} \cdot \frac{1}{6} = \frac{1}{6-5} = 1.$$

Thus  $P(X=\infty) = 1 - P(X \in \mathbb{N}) = 1 - 1 = 0$ .

Q. Does it mean  $\infty$  cannot happen?

## Discrete Random Variables

Random variables that take on at most a countable number of possible values are said to be discrete.

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If  $x_1, x_2, \dots$ , are the values that  $X$  can take with non zero probability, then  $p(x_n) \geq 0$  for  $n=1, 2, \dots$ , is called probability mass function pmf

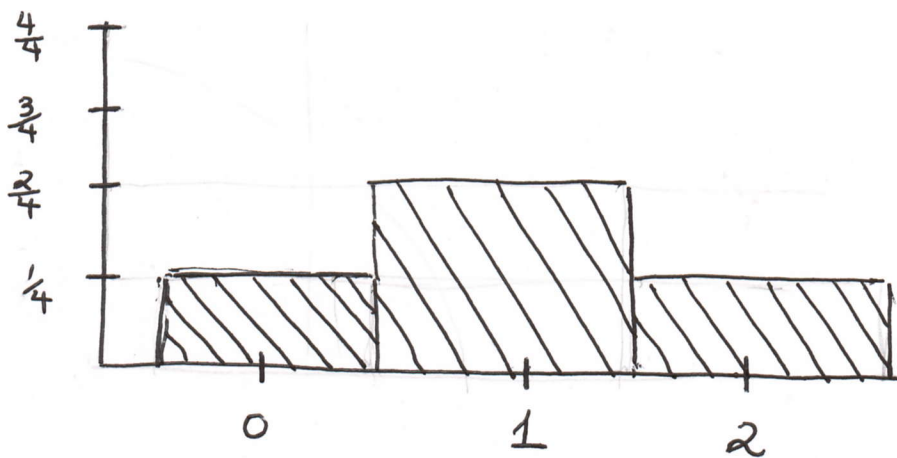
Since  $X$  must take on one of the values  $x_n$ , we

have 
$$\sum_{n=1}^{\infty} p(x_n) = 1.$$

Ex. Fair coin is tossed 2 times. Let  $X$  be the number of heads obtained. Then  $X = 0, 1, 2$ .

$$p(0) = p(\{TT\}) = \frac{1}{4}, \quad p(1) = p(\{TH, HT\}) = \frac{1}{2}.$$

$$p(2) = p(\{HH\}) = \frac{1}{4}.$$



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Ex. The probability mass function of a random variable  $X$  is given by  $p(n) = c \frac{\lambda^n}{n!}$   $n=0,1,2,\dots$

where  $\lambda$  is some positive value. Find

(a)  $P(X=0)$  and

(b)  $P(X>2)$

Solution: Since  $X$  must take one of the values

$0,1,2,\dots,$

$$1 = \sum_{n=0}^{\infty} c \frac{\lambda^n}{n!} = c \sum_{n=0}^{\infty} \frac{\lambda^n}{n!} = c e^{\lambda}. \text{ Thus } c = e^{-\lambda}$$

(a)  $P(X=0) = e^{-\lambda} \frac{\lambda^0}{0!} = e^{-\lambda}.$

(b)  $P(X>2) = 1 - P(X \leq 2) = 1 - (P(X=0) + P(X=1) + P(X=2))$   
 $= 1 - \left( e^{-\lambda} + e^{-\lambda} \frac{\lambda}{1} + e^{-\lambda} \frac{\lambda^2}{2} \right)$   
 $= 1 - e^{-\lambda} \left( 1 + \lambda + \frac{\lambda^2}{2} \right)$

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Cumulative Distribution

Let  $X$  be a discrete random variable and  $a \in \mathbb{R}$ .  
 We are often interested in problems of the form  
 $P(X \leq a) = F_X(a)$ .

If  $X$  takes on the values  $x_1, x_2, \dots$  then

$$F_X(a) = \sum_{n: x_n \leq a} P(x_n)$$

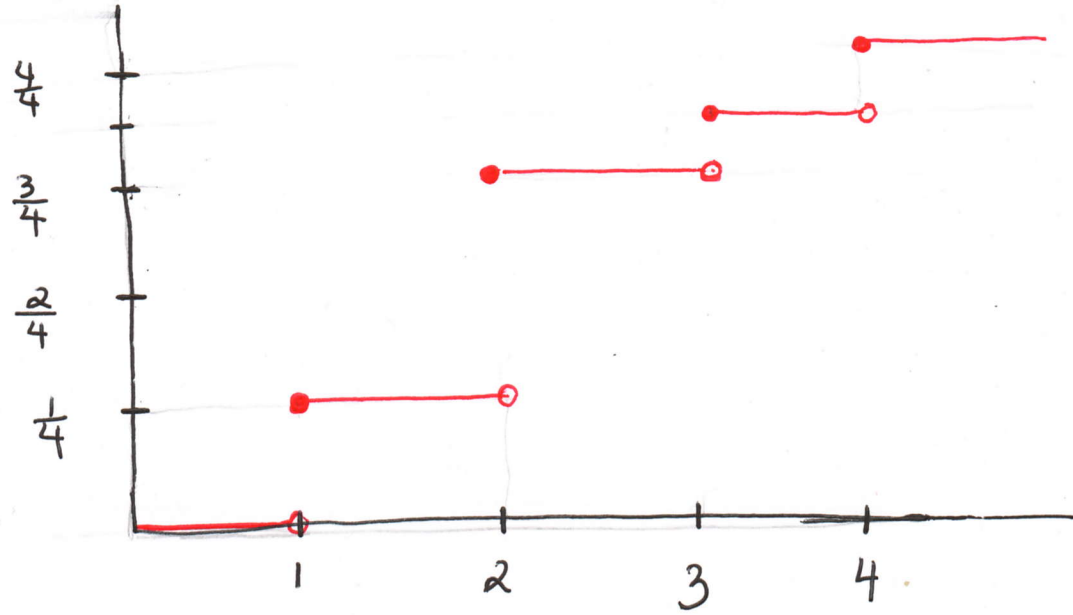
Ex. Suppose  $X$  has probability mass function

$$P(1) = \frac{1}{4}, \quad P(2) = \frac{1}{2}, \quad P(3) = \frac{1}{8}, \quad P(4) = \frac{1}{8}$$

Then

$$F_X(a) = \begin{cases} 0 & \text{if } a < 1 \\ \frac{1}{4} & \text{if } 1 \leq a < 2 \\ \frac{1}{4} + \frac{1}{2} & \text{if } 2 \leq a < 3 \\ \frac{1}{4} + \frac{1}{2} + \frac{1}{8} & \text{if } 3 \leq a < 4 \\ \frac{1}{4} + \frac{1}{2} + \frac{1}{8} + \frac{1}{8} & \text{if } 4 \leq a \end{cases}$$

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Cumulative distributions will become important when we speak about continuous random variables.

## Expected Value

Q. What happens when we play a game of chance habitually?

• If a game of chance has payoffs  $x_1, \dots, x_k$  with probabilities  $p(x_1), \dots, p(x_k)$ , what would be the average payoff per game if a large number of games,  $n$ , is played?

A.  $X = x_1, \dots, x_k$

$$\frac{x_1 n(X=x_1) + x_2 n(X=x_2) + \dots + x_k n(X=x_k)}{n}$$

$$= \sum_{j=1}^k \frac{x_j n(X=x_j)}{n} \quad \text{where } n(X=x_j) \text{ is the}$$

number of games among  $n$  that resulted in payoff  $x_j$ .

$$\text{Then } \lim_{n \rightarrow \infty} \sum_{j=1}^k x_j \frac{n(X=x_j)}{n} = \sum_{j=1}^k x_j \lim_{n \rightarrow \infty} \frac{n(X=x_j)}{n}$$

which equals...



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$$= \sum_{j=1}^k x_j p(x_j)$$

Def: Let  $X$  be a discrete random variable with values  $x_1, x_2, \dots$ . Then the expected value of  $X$ ,

$$E[X] = \sum_{j=1}^{\infty} x_j p(x_j)$$

$$= \sum_{x: p(x) > 0} x p(x)$$

This represents the weighted average of the values of  $X$ .

Ex. Find  $E[X]$ , where  $X$  is the outcome when we roll a fair die?

Solution: 
$$E[X] = 1 \cdot \frac{1}{6} + 2 \cdot \frac{1}{6} + 3 \cdot \frac{1}{6} + 4 \cdot \frac{1}{6} + 5 \cdot \frac{1}{6}$$

$$= \frac{1}{6} \sum_{k=1}^6 k = \frac{1}{6} \cdot \frac{6 \cdot 7}{2} = \frac{7}{2}$$

Ex. Suppose that when a fair die is rolled, you win \$6 when die comes up 6 and lose \$1 if the die comes up with another number.

Would you like to play this game? How about 1000,000 games?

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Solution: Let  $y = -1, 6$  be the possible payoffs.

$$E[Y] = -1 \cdot \frac{5}{6} + 6 \cdot \frac{1}{6} = 1 - \frac{5}{6} = \frac{1}{6}$$

Thus, if you play many games, the outcome will be the same as if you get  $\$ \frac{1}{6}$  per game.

If you play 1,000,000 games, your total payoff will be  $\$ \frac{1,000,000}{6}$ .

Ex. A game of Russian roulette is played with a revolver with 7 chambers. How many rounds do we expect the game to last?

Solution: We can solve this problem with just a little intuition.

The bullet is in one chamber, so the pistol fires once every...

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yes exactly! Every 7 rounds! Thus we expect the gun to fire once every 7 rounds on the average, making the tournament's duration typically 7 rounds long.

Lets solve this rigorously.

$X = 1, 2, 3, \dots$  be the number of rounds.

$$P(X=n) = \underbrace{\left(\frac{6}{7}\right)^{n-1}}_{n-1 \text{ misfires}} \underbrace{\frac{1}{7}}_{\text{fires on the last round.}}$$

$$E[X] = \sum_{n=1}^{\infty} n \left(\frac{6}{7}\right)^{n-1} \frac{1}{7} \quad \text{This series reminds me of the movie Deer Hunter.}$$

Recall:  $G = \sum_{n=1}^{\infty} x^{n-1} = 1 + x + x^2 + \dots +$

$$= 1 + x \underbrace{(1 + x + \dots +)}_G = 1 + xG \quad ; \quad G = \frac{1}{1-x}$$

$$H = \sum_{n=1}^{\infty} n x^{n-1} = \sum_{n=1}^{\infty} [(n-1)x^{n-1} + x^{n-1}] = \sum_{n=1}^{\infty} (n-1)x^{n-1} + \frac{1}{1-x}$$

$$= \sum_{m=1}^{\infty} m x^m + \frac{1}{1-x} = x \sum_{m=1}^{\infty} m x^{m-1} + \frac{1}{1-x} = xH + \frac{1}{1-x}$$

$$\text{So } H = \frac{1}{(1-x)^2} \quad \text{in particular } E[X] = \frac{1}{\left(1-\frac{6}{7}\right)^2} \cdot \frac{1}{7}$$
$$= \frac{1}{\left(\frac{1}{7}\right)^2} \cdot \frac{1}{7} = \frac{1}{\left(\frac{1}{7}\right)} = 7.$$